



Published in final edited form as:

AIDS Res Hum Retroviruses. 2017 June ; 33(6): 546–554. doi:10.1089/AID.2016.0203.

Estimating False-Recent Classification for the Limiting-Antigen Avidity EIA and BED-Capture Enzyme Immunoassay in Vietnam: Implications for HIV-1 Incidence Estimates

Neha S. Shah¹, Yen T. Duong¹, Linh-Vi Le², Nguyen Anh Tuan³, Bharat S. Parekh¹, Hoang Thi Thanh Ha³, Quang Duy Pham⁴, Cao Thi Thu Cuc⁴, Trudy Dobbs¹, Tran Hong Tram³, Truong Thi Xuan Lien⁴, Nick Wagar¹, Chunfu Yang¹, Amy Martin¹, Mitchell Wolfe², Nguyen Tran Hien³, and Andrea A. Kim¹

¹Centers for Disease Control and Prevention, Atlanta, Georgia

²Centers for Disease Control and Prevention, Hanoi, Vietnam

³National Institute of Hygiene and Epidemiology, Hanoi, Vietnam

⁴Ho Chi Minh City Pasteur Institute, Ho Chi Minh City, Vietnam

Abstract

Laboratory tests that can distinguish recent from long-term HIV infection are used to estimate HIV incidence in a population, but can potentially misclassify a proportion of long-term HIV infections as recent. Correct application of an assay requires determination of the proportion false recent (PFRs) as part of the assay characterization and for calculating HIV incidence in a local population using a HIV incidence assay. From April 2009 to December 2010, blood specimens were collected from HIV-infected individuals attending nine outpatient clinics (OPCs) in Vietnam (four from northern and five from southern Vietnam). Participants were living with HIV for 1 year and reported no antiretroviral (ARV) drug treatment. Basic demographic data and clinical information were collected. Specimens were tested with the BED capture enzyme immunoassay (BED-CEIA) and the Limiting-antigen (LAg)-Avidity EIA. PFR was estimated by dividing the number of specimens classified as recent by the total number of specimens; 95% confidence intervals (CI) were calculated. Specimens that tested recent had viral load testing performed. Among 1,813 specimens (north, $n = 942$ and south, $n = 871$), the LAg-Avidity EIA PFR was 1.7% (CI: 1.2–2.4) and differed by region [north 2.7% (CI: 1.8–3.9) versus south 0.7% (CI: 0.3–1.5); $p = .002$]. The BED-CEIA PFR was 2.3% (CI: 1.7–3.0) and varied by region [north 3.4% (CI: 2.4–4.7) versus south 1.0% (CI: 0.5–1.2), $p < .001$]. Excluding specimens with an undetectable VL, the LAg-Avidity EIA PFR was 1.2% (CI: 0.8–1.9) and the BED-CEIA PFR was 1.7% (CI: 1.2–2.4). The LAg-Avidity EIA PFR was lower than the BED-CEIA PFR. After excluding specimens with

Address correspondence to: Neha S. Shah, Centers for Disease Control and Prevention, 850 Marina Bay Parkway, Building P, 2nd Floor, Richmond, CA 94804, nshah6@cdc.gov.

Disclaimer

The findings and conclusions in this article are those of the authors and do not necessarily represent the official position of the Centers for Disease Control and Prevention.

Author Disclosure Statement

No competing financial interests exist.

an undetectable VL, the PFR for both assays was similar. A low PFR should facilitate the implementation of the LAg-Avidity EIA for cross-sectional incidence estimates in Vietnam.

Keywords

HIV; prevention; epidemiology; incidence

Introduction

The first case of HIV infection in Vietnam was reported in 1990. An estimated 256,000 people were living with HIV in Vietnam in 2014.¹ Currently the prevalence of HIV infection in Vietnam is less than 1% and concentrated within key populations at risk, including injection drug users (IDU), female sex workers (FSW), and men who have sex with men (MSM). Between 2000 and 2014, the estimated number of people living with HIV more than doubled from 100,000 to 256,000.^{1,2} Although significant progress has been made in responding to the HIV epidemic in the country, international donor contributions are decreasing and much work is still needed to reach Vietnam's goal of ending AIDS by 2030.

Identifying subpopulations at high risk for new HIV infection is critical for halting the transmission of HIV. Currently most HIV programs worldwide track the epidemic by monitoring trends in HIV prevalence in a population. Although prevalence is an important measure, it provides a limited understanding of the most recent spread of infection or HIV incidence. HIV incidence can identify current transmission dynamics to effectively target interventions and allocate resources. In addition, HIV incidence provides data for evaluating the success of prevention programs.³ However, obtaining HIV incidence rates have been challenging. Longitudinal prospective cohort studies are the gold standard for estimating HIV incidence, yet they are lengthy, costly, complicated, and may be unrepresentative of the general population due to self-selective enrollment of participants, among other issues.³ In response, laboratory assays were developed that can distinguish recent from long-term HIV infection based on immunologic markers of disease progression and can be applied to cross-sectional surveys to estimate HIV incidence.⁴ The World Health Organization's (WHO) guidelines for estimating HIV incidence in the population have incorporated the use of one or more tests as part of an algorithm to estimate cross-sectional HIV incidence.³

A significant limitation of current HIV incidence laboratory assays is that they overestimate HIV-1 incidence by misclassifying a proportion of individuals with long-term infection (defined in this study as infection ≥ 1 year) as recent infection, which can lead to errors in incidence estimations.^{4,5} The proportion of individuals with long-term infection that misclassify as recent infection on the incidence assay is termed in this study as the assay's proportion false recent (PFR). Because the PFR has been shown to vary significantly by population, HIV subtype,⁶ HIV epidemic phases,⁷ individual immune status, and antiretroviral (ARV) use,^{8–10} current guidance for estimating incidence recommends that countries determine a local PFR that can be incorporated into the incidence formula for calculating HIV incidence.^{3,11} Low viral loads among HIV-infected persons as a result of ARV treatment and among those who naturally maintain an undetectable viral load in the

absence of treatment (also known as “elite controllers”) will misclassify as false recent on incidence assays that rely on HIV antibodies to determine recent infection. Experts in the field of HIV incidence assay development recommend that a PFR for a HIV incidence assay not exceed 2% to produce reliable incidence estimates.⁶

At the time of this study, the BED capture enzyme immunoassay (BED-CEIA) was the only commercially available assay, and the Limiting-antigen (LAG)-Avidity EIA was in development. However, the new single-well LAG-Avidity EIA was recently described to distinguish recent from longstanding HIV infections for the purpose of estimating incidence in cross-sectional populations.¹² Briefly, the BED-CEIA measures the proportion of anti-HIV IgG relative to total IgG. As anti-HIV IgG increases with time, a lower proportion of anti-HIV IgG to total IgG would indicate a recent infection.^{13,14} In contrast, the LAG-Avidity EIA determines the avidity or “binding strength” of the HIV-1 antibodies. Because antibodies mature and become more specific over time, antibodies from a recently infected individual would have a weaker binding strength to the virus than an individual infected for a longer duration.^{12,15} While the BED-CEIA has shown relatively high PFR in a standardized multiclade specimen set derived from persons with longstanding HIV infection, the PFR for the LAG-Avidity EIA has been shown to be substantially lower in the same samples.¹⁶ However, subtype-specific PFR for clade A/E, which predominates in South East Asia, was not included in this evaluation. The purpose of this study was to determine the applicability of the LAG-Avidity EIA and the BED-CEIA in Vietnam by determining the PFR for each assay.

Methods

Study population

This cross-sectional study was conducted from April 2009 to December 2010 in nine purposively selected outpatient clinics (OPCs) in the northern ($n = 4$) and southern regions ($n = 5$) of Vietnam. Eligible participants were 18 years of age or older, had a confirmed HIV infection for 1 year based on initial HIV diagnosis date, and self-reported no ARV use. A sample size of 1,917 specimens was calculated using the BED-CEIA parameters and assuming an expected PFR of 5% (based on previously published studies), measured with precision of 1% (precision was defined as half the width of the confidence interval), and a refusal rate of 5%.^{13,17–19}

Specimen collection

After informed consent was provided, a venous blood sample was collected from each individual, and a chart review was conducted to obtain data on demographics, HIV risk factors, opportunistic infections (OIs), and CD4 count. CD4 count was based on the most recent CD4 result documented in the medical chart at the time of enrollment. The whole blood was processed into plasma aliquots by centrifuging the sample at 3,000 rpm for 10min and stored at -70°C .

Laboratory testing for recent HIV-1 infection

Specimens were tested with the LAg-Avidity EIA and BED-CEIA according to the manufacturer's instructions (both manufactured by Sedia BioSciences, Portland, Oregon). Normalized optical density (ODn) cutoffs of 1.5 for LAg-Avidity EIA and 0.8 for BED-CEIA were used to distinguish recent from long-term HIV infection. Specimens with final ODn values at or below the ODn cutoff for each respective assay were classified as recent infections, while those with values above the ODn cutoff were classified as long-term infections. BED-CEIA testing was conducted by the Vietnam National Institute of Hygiene and Epidemiology (NIHE) and the Ho Chi Minh City Pasteur Institute (HCMC PI). The LAg-Avidity EIA testing was performed at the Centers for Disease Control and Prevention (CDC) in Atlanta, Georgia, where specimens were tested in duplicate runs to increase confidence of results irrespective of their ODn levels. The mean ODn value was used for final classification. All specimens classified as recent with ODn values lower than 0.4 and 0.3 by the LAg-Avidity EIA and BED-CEIA, respectively, were further tested by western blot to confirm HIV seropositivity at the CDC. All specimens that classified as recent and that had sufficient volume underwent additional testing, including HIV subtyping, viral load analysis, and detection of ARV drugs at the CDC.

HIV-1 subtyping

Subtyping was performed on a random sample of 100 specimens. Nucleic acid was isolated from 200 µL plasma samples using the NucliSENS® system for lysis and extraction (BioMerieux). Nucleic acid was eluted in 25 µL of NucliSENS Extraction Buffer 3 and either immediately used for RT-PCR or stored at -80°C until use. Subtyping of the HIV-1 *env*gp41 gene was performed following the procedure of a broadly sensitive CDC in-house assay as described in detail previously with the exception of using *env*gp41 primers as described before.^{20,21} The ReCALL software program was used to edit the raw sequences and generate consensus sequences.²²

HIV-1 subtyping for the newly obtained sequences was performed using the REGA HIV-1 Genotyping Tool.²³ Phylogenetic analyses were further conducted using neighbor-joining method included in the MEGA 5 for sequences with unclassifiable subtypes.²³ Reference sequences were obtained from the Los Alamos HIV Database (www.hiv.lanl.gov). The stability of the tree nodes was assessed by bootstrap analysis using 1,000 replicates. Bootstrap values > 70% were considered significant.²⁴

Viral load testing

Quantification of plasma RNA was determined using the Abbott RealTime m2000 platform. Due to low sample volumes, plasma samples were diluted fourfold in serum base matrix, resulting in a limit of detection of 600 copies/mL. Specimens with viral load <1,000 copies/mL were considered to have a low viral load.

Detection of ARV drugs

Specimens that classified as recent infection by either assay were tested for the detection of antiretroviral drugs to ensure that individuals were indeed ARV-naïve. Antiretroviral (ARV) drugs from first-line regimens in Vietnam [Nevirapine (NVP), Efavirenz (EFV), and

Lamivudine (3TC)] were measured in plasma simultaneously by high-performance liquid chromatography– tandem mass spectrometry (HPLC-MS/MS). Drugs from 100 µL of plasma were extracted with 370 µL of acetonitrile containing 270 ng/mL each of the internal standards [NVP-d5 (TRC) for NVP, EFV-d4 (TRC) for EFV, and FTC-13C 15 N2 (Moravek Biochemicals and Radiochemicals) for 3TC]. Protein precipitates were removed by centrifugation, and the liquid extracts were transferred to the wells of a 96-well polypropylene plate and then evaporated to almost dryness in a vacuum concentrator and reconstituted with 150 µL of mobile phase A (0.1% formic acid in water). A volume of 15 µL of the final processed sample was injected into a Shimadzu Prominence HPLC system, connected to a model 3200 QTRAP mass spectrometer (AB Sciex) that was used to detect the compounds of interest. Data were processed using Analyst 1.5.1 (AB Sciex). The limit of quantification (LOQ) was 25 ng/mL for NVP and 10 ng/mL for both EFV and 3TC.

Calculation of PFR

The PFR was estimated by dividing the number of specimens testing recent on the assay by the total number of specimens in the study population. The PFR was calculated for each assay by select demographic and clinical characteristics and stratified by geographic region (northern versus southern) and clinic. We used the term PFR instead of false recent rate (FRR) as the latter implies a time factor which was not addressed in this study. Statistical significance was assessed using *t*-test, chi-square, McNemar, and Fisher's exact tests, where indicated. *p*-values of <.05 were considered statistically significant. Confidence intervals were calculated using the Wilson score interval procedure. Multivariate analysis was not conducted as the study was not powered for such analysis. Data forms were double entered into EpiData by NIHE and HCMC PI. Data analyses were conducted using Stata version 10.0 (StataCorp) and SAS version 9.2 (SAS Institute).

Ethical considerations

The protocol was reviewed and approved by the Institutional Review Boards of CDC and NIHE. Because these incidence assays may not be accurate at the individual level and are approved for surveillance use only, individual-level data were not linked to any identifiers and test results were not returned to participants. Specimens and data forms were labeled with unique study identification (ID) numbers, which were used to merge clinical data and incidence assay results.

Results

Characteristics of the study population

A total of 1,935 samples were collected (north, $n = 1,030$; south, $n = 905$). After excluding specimens with missing data, duplicates, and from individuals not meeting the inclusion criteria, including those who were classified as recent but were positive for ARV drugs, the final sample size for analysis was 1,813.

Table 1 shows the demographic, clinical, and risk factor characteristics of the population enrolled in the study. There were regional differences in the age distribution (north: 30.6% were aged ≥ 35 years versus south: 16.2%); CD4 cell count (north: 28.5% had CD4 count

<200 cells/mm³ compared to south: 15.8%); and duration of HIV infection (north: 30.9% had documented HIV infection >3 years compared to south: 11.5%). There were limited data on OIs; however, the most common documented OI was TB (11.2% of the study population).

Determination of PFRs

Overall, 31 of 1,813 specimens (1.7%; CI: 1.2–2.4) were classified as recent by the LAg-Avidity EIA (Table 2), and 41 (2.3%; CI: 1.7–3.1) were classified as recent by the BED-CEIA. Twenty-four (77.4%) of 31 LAg-recent specimens were also recent by the BED-CEIA; seven specimens testing recent on the LAg-Avidity EIA did not test as recent on the BED-CEIA. All specimens classified as false recent by either LAg-Avidity EIA or BED-CEIA were recombinant subtype CRFO1_AE. The LAg-Avidity EIA or BED-CEIA PFRs were not statistically significantly different (p -value .06).

Analyses of PFRs

Table 3 summarizes the BED-CEIA PFR, by selected demographic and clinical characteristics, stratified by region. Thirty-two of 942 (3.4%, CI: 2.4–4.7) specimens from the north were false recent compared to 9 of 873 (1.0%, CI: 0.5–1.2) specimens from the south ($p < .001$). Within northern specimens, the BED-CEIA PFR was similar between sex (3.4% and 3.5% for specimens from men and women, respectively), while within southern specimens, more specimens from women (1.7%) were classified as false recent compared to those from men (0.2%). The differences in PFR by clinical characteristics and risk groups were not significant, and the differences in PFR between OIs were not analyzed due to zero cells.

Table 4 summarizes the PFR for LAg-Avidity EIA and by select demographic and clinical characteristics, stratified by region. Twenty-five of 942 (2.7%, CI: 1.8–3.9) specimens from the north misclassified as recent compared to 6 of 871 (0.7%, CI: 0.3–1.5) from the south ($p = .002$). There were no statistically significant differences in demographic and clinical characteristics between regions. However, specimens from individuals living with HIV for 1–2 years had a higher proportion of misclassified specimens (2.0% vs. 1.5% among those infected >2 years).

For both assays, the PFR differed significantly by region with northern clinics having a higher PFR on both assays (Table 2). The PFR for both assays differed significantly among the northern clinics ($p < .05$, Fig. 1). The BED-CEIA PFR for northern clinics B (5.8%; CI: 3.2–10.3) and C (4.2%; CI: 2.6–6.7) was high compared to northern clinics A (1.5%; CI: 0.6–3.8) and D (1.7%; CI: 0.5–5.8). Specimens from northern clinic B (6.3%; CI: 3.6–11.0) had a higher LAg-Avidity EIA PFR compared to northern clinics A (0.4%; CI: 0.1–2.1), C (3.1%; CI: 1.8–5.4), and D (0.8%; CI: 0.2–4.5). The southern clinics had a similar PFR for both the LAg-Avidity EIA and BED-CEIA ($p > .1$).

Residual PFR

Among false-recent specimens, 10 specimens on BED-CEIA and 9 specimens on LAg-Avidity EIA had HIV-1 RNA <1,000 copies/mL (low viral load, Fig. 2). After excluding

specimens with low viral load, the residual PFR for BED-CEIA was 1.6% (CI: 1.2–2.4) and for LAg-Avidity EIA was 1.2% (CI: 0.8–1.9) and not significantly different. Regionally, the residual BED-CEIA PFR was 2.0% (CI: 1.3–3.2) for the north and 0.3% (CI: 0.1–1.0) for the south. Similarly, the residual LAg-Avidity EIA PFR was 2.8% (CI: 1.9–4.0) for the north and 0.6% (CI: 0.3–1.4) for the south.

Discussion

This is the first study to estimate the PFRs in Southeast Asia for both the LAg-Avidity EIA and BED-CEIA from the same population. Although there was no statistically significant difference between the two assays, the point estimates for the LAg-Avidity EIA PFR were lower than the BED-CEIA, which complements previous demonstrations of improved performance characteristics of the LAg-Avidity EIA compared with the BED-CEIA.^{16,25} The better performance of LAg-Avidity EIA is likely attributed to the design of the assay to measure HIV antibody avidity compared to the BED-CEIA, which is a capture EIA that measures the proportion of anti-HIV IgG to total IgG.

After accounting for low viral loads, the PFR decreased to 1.2% for the LAg-Avidity EIA and 1.7% for the BED-CEIA. Both the LAg-Avidity EIA and the BED-CEIA PFR were consistent with recent literature.^{26,27} The BED-CEIA PFR is the same as was reported from samples collected only from HCMC in southern Vietnam, however, was lower than previously reported in other populations.^{27–30} The LAg-Avidity EIA was lower than the AxSYM avidity index assay FRR of 2.7% from the HCMC study.²⁷ The difference in PFR after excluding individuals with low viral loads demonstrates how individuals with long-term infection who are elite controllers with naturally low or undetectable viral loads or potential individuals on ARV therapy who have low antibody levels can misclassify as recent on antibody-based assays.^{5,31,32} Therefore, to obtain accurate incidence estimates using current incidence assays, misclassification from low viral loads should be expected and accounted for in a recent infection testing algorithm. Although a residual PFR may still be present after these adjustments, in the context of Vietnam, given that the level of the PFR fell within the WHO recommended threshold (<2%), further adjustments may not be needed.

For example, we demonstrate the impact of the LAg-Avidity PFR on HIV incidence estimates in a key population at high risk for HIV exposure in Vietnam, persons who inject drugs (PWID) using a hypothetical example based on epidemiological data in this population.³³ For this hypothetical example, the MDRI estimates were not adjusted to account for viral load. If we assume that in a sample of 1,000 PWID from the Northern region of the country, 11% are HIV positive and 10% of these are recent on the LAg Avidity, HIV incidence would be 2.4% (CI: 0.9–3.8) using no PFR, 1.9% (CI: 0.4–3.5) using our observed PFR 2.7%, and 2.1% (CI: 0.6–3.6) using our observed residual PFR of 2.0%. Similar calculations for the BED-CEIA would yield an incidence of 2.2% (CI: 0.8–3.6) using no PFR and 1.5% (CI: 0.2–2.9) using our observed PFR 3.4% and 1.7% (CI: 0.3–3.0). The minimal differences observed between the incidence estimates after further exclusion of low viral loads in our study and in the above example suggest that for this hypothetical PWID population example in Vietnam, the inclusion of viral load in the recent infection testing algorithm may not be warranted. Finally, similar to other evaluations, we found that

specimens classified as false recent by the LAg-Avidity EIA assay were also false recent on the BED-CEIA^{12,15} and a two-test algorithm using the BED-CEIA as the screening assay followed by LAg-Avidity EIA as the confirmatory assay would result in a PFR of 2.2%, which is not significantly different compared to LAg-Avidity EIA alone.

There were statistically significant regional differences in the BED-CEIA and LAg-Avidity EIA PFR between north and south Vietnam and differences by northern clinic site. The differences in age, CD4 count, and duration of HIV infection in the north compared to the south confirm regional differences between the populations of HIV-infected people in care. All false-recent cases were HIV-1 subtype AE, the most dominant subtype in the region, confirming that HIV-1 clade did not contribute to observed regional differences in the PFR. However, there were differences in the PFRs by clinic site suggesting differing populations by OPCs.

This study was subject to several limitations. The majority of the misclassified specimens were from the two northern clinics. In combination, this suggests there may be clinic-specific biases related to data quality from these clinics. This study was not powered to detect differences in the PFR by site of enrollment and demographic and clinical characteristics in the northern and southern regions separately. Furthermore, determining a site-specific PFR is not feasible given the large sample size needed to determine a PFR. Participants were recruited at the point of care at the HIV clinic, and therefore, results are not generalizable to the entire population of persons living with HIV in Vietnam. Furthermore, the southern region was represented only by one province, Ho Chi Minh City, whereas in the northern region, samples were collected from three provinces. Although no population level samples are available to determine a more generalizable PFR, our results are consistent with results from a study in China showing a LAg-Avidity EIA FRR of 1.1% and HCMC showing a BED-CEIA FRR of 1.7%. Finally, since data on OIs were limited, we were unable to assess possible associations between OI and false-recent classification that may be associated with false-recent misclassification due to advanced HIV status and decreased antibody response.

We are continuing to examine the application of the LAg-Avidity EIA in cross-sectional population in Vietnam and elsewhere to assess its utility in providing accurate incidence estimates and trends over time. Use of the PFR may result in an over, under, or even a negative estimate of incidence if the PFR is not derived from a population similar to the one that is being surveyed for incidence estimation or if the PFR is elevated due to sample bias (e.g., ARV use). Therefore, caution should be taken during study design for HIV incidence estimation and sample collection to minimize potential biases. As suggested by the WHO recommendations, a multiassay algorithm that incorporates testing for viral load at a minimum, and testing for ARV drugs where feasible, should be conducted to minimize false recency.³⁴ The data presented in this study are consistent with a previously described low PFR for the LAg-Avidity EIA and demonstrate a low PFR for the BED-CEIA.¹² Therefore, we believe that both assays have the potential to provide accurate HIV-1 incidence in cross-sectional populations in Vietnam. However, when ARV and viral load testing are not available, because the PFR was slightly lower with the LAg-Avidity EIA (1.7% versus 2.3% with the BED-CEIA without VL testing), the LAg-Avidity EIA should be used for

cross-sectional incidence assays and for impact evaluation of combination prevention programs being implemented to reduce the transmission of HIV in Vietnam. Finally, at this time we do not know if the PFR is stable over time and more research is needed to better understand if there is variability to the PFR as HIV epidemic change.

Acknowledgments

The authors thank all those who participated in study and the local organizations that assisted with study enrollment and data collection.

References

1. Viet Nam Ministry of Health. Optimizing Viet Nam's HIV Response: An Investment Case. Viet Nam Ministry of Health; Hanoi: 2014.
2. Viet Nam Ministry of Health. Vietnam HIV/AIDS Estimates and Projections 2007–2012. Vietnam Administration of HIV/ AIDS Control. Family Health International; Hanoi: 2009.
3. World Health Organization. When and How to Use Assays for Recent Infection to Estimate HIV Incidence at a Population Level. WHO technical working group on HIV incidence assays; Geneva: 2011.
4. Mastro TD, Kim AA, Hallett T, et al. Estimating HIV incidence in populations using tests for recent infection: Issues, challenges and the way forward. *J HIV/AIDS Surveill Epidemiol*. 2010; 2:1–14.
5. Laeyendecker O, Brookmeyer R, Oliver AE, et al. Factors associated with incorrect identification of recent HIV infection using the BED capture immunoassay. *AIDS Res Hum Retroviruses*. 2012; 28:816–822. [PubMed: 22014036]
6. Incidence Assay Critical Path Working Group. More and better information to tackle HIV epidemics: Towards improved HIV incidence assays. *PLoS Med*. 2011; 8:e1001045. [PubMed: 21731474]
7. Hallett TB, Ghys P, Barnighausen T, Yan P, Garnett GP. Errors in 'BED'-derived estimates of HIV incidence will vary by place, time and age. *PLoS ONE*. 2009; 4:e5720. [PubMed: 19479050]
8. Laeyendecker O, Rothman R, Henson C, et al. The effect of viral suppression on cross-sectional incidence testing in the Johns Hopkins Hospital Emergency Department. *J Acquir Immune Defic Syndr*. 2008; 48:211–215. [PubMed: 18520680]
9. Hladik W, Olara D, Mermin J, et al. Effect of CD4(+) T Cell count and antiretroviral treatment on two serological HIV incidence assays. *AIDS Res Hum Retroviruses*. 2012; 28:95–99. [PubMed: 21314476]
10. Hayashidam T, Gatanaga H, Tanuma J, Oka S. Effects of low HIV type 1 load and antiretroviral treatment on IgG-capture BED-enzyme immunoassay. *AIDS Res Hum Retroviruses*. 2008; 24:495–498. [PubMed: 18327979]
11. Office of the Global AIDS Coordinator. Statement from the Surveillance and Survey and the Laboratory Working Groups to the Office of the Global AIDS Coordinator. OGAC; Washington: DC: 2010. Update on HIV-1 Incidence estimation and surveillance in resource-constrained settings using tests for recent infection.
12. Duong YT, Qiu M, De AK, et al. Detection of recent HIV-1 infection using a new limiting-antigen avidity assay: Potential for HIV-1 incidence estimates and avidity maturation studies. *PLoS ONE*. 2012; 7:e33328. [PubMed: 22479384]
13. Hargrove JW, Humphrey JH, Mutasa K, et al. Improved HIV-1 incidence estimates using the BED capture enzyme immunoassay. *AIDS*. 2008; 22:511–518. [PubMed: 18301064]
14. Parekh BS, Kennedy MS, Dobbs T. Quantitative detection of increasing HIV Type 1 antibodies after seroconversion: A simple assay for detecting recent HIV infection and estimating incidence. *AIDS Res Hum Retroviruses*. 2002; 18:295–307. [PubMed: 11860677]
15. Kim, A., Parekh, B., Muro, M., et al. Validation of the Limiting Antigen Avidity Enzyme Immunoassay as a Tool to Estimate National HIV Incidence; 2007 AIDS Indicator Survey: Kenya. Conference on Retroviruses and Opportunistic Infections; Seattle. 2012;

16. Kassanjee R, Pilcher CD, Keating SM, et al. Independent assessment of candidate HIV incidence assays on specimens in the CEPHIA repository. *AIDS* (London, England). 2014; 28:2439–2449.
17. McDougal JS, Parekh BS, Peterson M, Branson B, Dobbs T, Ackers M, Gurwith M. Comparison of HIV Type 1 incidence observed during lLongitudinal follow-up with incidence estimated by cross-sectional analysis using the BED capture enzyme immunoassay. *AIDS Res Hum Retroviruses*. 2006; 22:945–952. [PubMed: 17067263]
18. Dobbs T, Kennedy S, Pau C-P, McDougal JS, Parekh BS. Performance characteristics of the immunoglobulin G-capture BED-enzyme immunoassay, an assay to detect recent human immunodeficiency virus type 1 seroconversion. *J Clin Microbiol*. 2004; 42:2623–2628. [PubMed: 15184443]
19. Kim A, McDougal J, Hargrove J, et al. Evaluating the BED capture enzyme immunoassay to estimate HIV incidence among adults in three countries in sub-Saharan Africa. *AIDS Res Hum Retroviruses*. 2010; 26:1051–1061. [PubMed: 20849299]
20. Yang C, Pieniazek D, Owen SM, et al. Detection of phylogenetically diverse human immunodeficiency virus type 1 groups M and O from plasma by using highly sensitive and specific generic primers. *J Clin Microbiol*. 1999; 37:2581–2586. [PubMed: 10405405]
21. Yang C, Dash BC, Simon F, et al. Detection of diverse variants of human immunodeficiency virus-1 groups M, N, and O and simian immunodeficiency viruses from chimpanzees by using generic pol and env primer pairs. *J Infect Dis*. 2000; 181:1791–1795. [PubMed: 10823786]
22. Woods CK, Brumme CJ, Liu TF, et al. Automating HIV drug resistance genotyping with RECall, a freely accessible sequence analysis tool. *J Clin Microbiol*. 2012; 50:1936–1942. [PubMed: 22403431]
23. de Oliveira T, Deforche K, Cassol S, et al. An automated genotyping system for analysis of HIV-1 and other microbial sequences. *Bioinformatics*. 2005; 21:3797–3800. [PubMed: 16076886]
24. Hall TA. BioEdit: A user-friendly biological sequence alignment editor and analysis program for Windows 95/98/NT. *Nucleic Acids Symposium Series*. 1999; 41:95–98.
25. Duong YT, Kassanjee R, Welte A, et al. Recalibration of the limiting antigen avidity EIA to determine mean duration of recent infection in divergent HIV-1 subtypes. *PLoS ONE*. 2015; 10
26. Yu L, Laeyendecker O, Wendel SK, et al. Low false recent rate of limiting-antigen avidity assay among long-term infected subjects from Guangxi, China. *AIDS Res Hum Retroviruses*. 2015; 31:1247–1249. [PubMed: 26331573]
27. Sexton CJ, Costenbader EC, Vinh DT, et al. Correlation of prospective and cross-sectional measures of HIV type 1 incidence in a higher-risk cohort in Ho Chi Minh City, Vietnam. *AIDS Res Hum Retroviruses*. 2012; 28:866–873. [PubMed: 21936716]
28. Wei X, Liu X, Dobbs T, et al. Development of two avidity-based assays to detect recent HIV type 1 seroconversion using a multisubtype gp41 recombinant protein. *AIDS Res Hum Retroviruses*. 2009; 26:61–71.
29. Bärnighausen T, McWalter TA, Rosner Z, Newell M-L, Welte A. HIV incidence estimation using the BED capture enzyme immunoassay: Systematic review and sensitivity analysis. *Epidemiology*. 2010; 21:685–69. [PubMed: 20699682]
30. Hallett TB. Estimating the HIV incidence rate: Recent and future developments. *Curr Opin HIV AIDS*. 2011; 6:102–107. [PubMed: 21505383]
31. Laeyendecker O, Brookmeyer R, Mullis CE, et al. Specificity of four laboratory approaches for cross-sectional HIV incidence determination: Analysis of samples from adults with known nonrecent HIV infection from five african countries. *AIDS Res Hum Retroviruses*. 2012; 28:1–7. [PubMed: 21495876]
32. Marinda ET, Hargrove J, Preiser W, et al. Significantly diminished long-term specificity of the BED capture enzyme immunoassay among patients with HIV-1 with very low CD4 counts and those on antiretroviral therapy. *J Acquir Immune Defic Syndr*. 2010; 53:496–499. [PubMed: 20306555]
33. Viet Nam Administration of AIDS Control. Vietnam AIDS Response Progress Report 2014. Ministry of Health; Ha Noi: 2014.

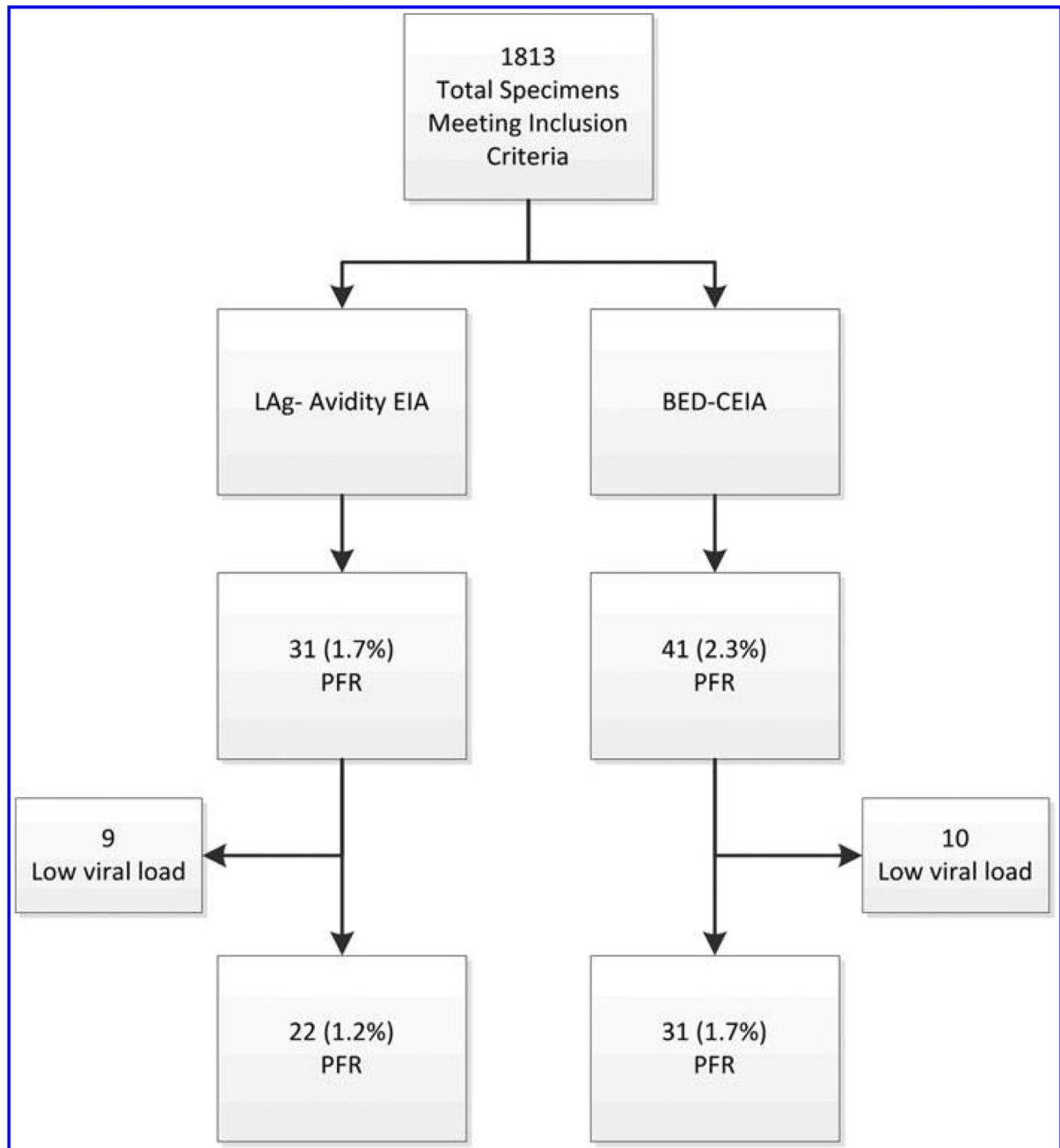
34. World Health Organization and UNAIDS. WHO HIV Incidence Assay Working Group Meeting. UNAIDS, Barcelona; Spain: 2015. Technical Update on HIV Incidence Assays for Surveillance and Monitoring purposes.

Author Manuscript

Author Manuscript

Author Manuscript

Author Manuscript

**FIG. 1.**

The PFR for the LAg-Avidity EIA and the BED-CEIA by clinic site, Vietnam 2009–2010.

PFR, proportion false recent; BED-CEIA, BED capture enzyme immunoassay.

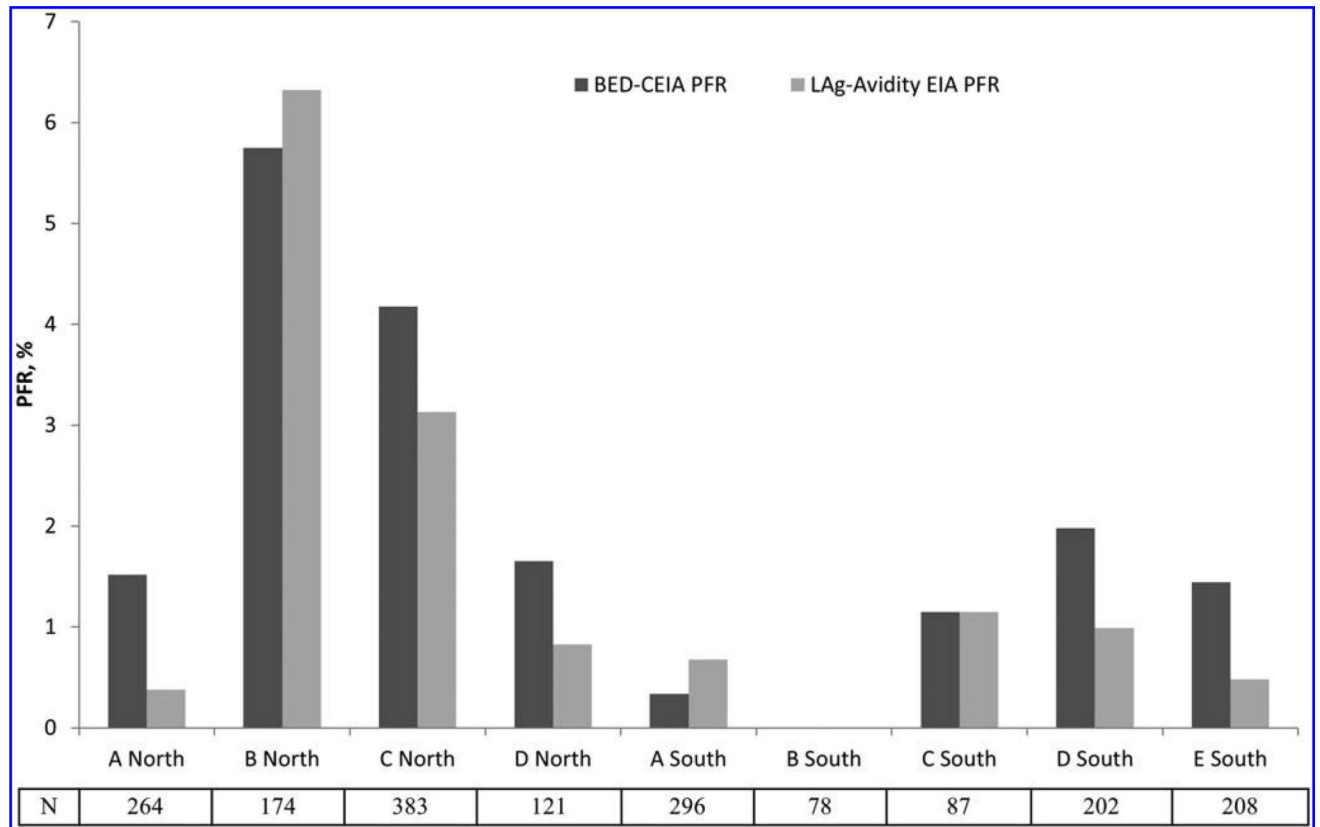


FIG. 2.
PFR algorithm for the LAg-Avidity EIA and BED-CEIA excluding ARV and ARV plus viral load, Vietnam 2009–2010. ARV, antiretroviral.

Table 1

Characteristics of PFR Study Sample Stratified by Region, Vietnam 2009–2010

Characteristic	Total (n, %) N = 1813	North (n, %) N = 942	South (n, %) N = 871	p-value
Age group (years)				<.001
18–24	201 (11.1)	68 (7.2)	133 (15.3)	
25–34	1,182 (65.2)	585 (62.1)	597 (68.5)	
35–44	355 (19.6)	234 (24.8)	121 (13.9)	
45–54	67 (3.7)	50 (5.3)	17 (2.0)	
55+	8 (0.4)	5 (0.5)	3 (0.3)	
Sex				<.001
Male	968 (53.7)	558 (59.8)	410 (47.1)	
Female	836 (46.3)	375 (40.2)	461 (52.9)	
Missing	9 (0.5)	9 (1.0)		
Duration since documented HIV positive test result (years)				<.001
1–2	853 (47.1)	408 (43.3)	445 (51.1)	
2–3	569 (31.2)	243 (25.8)	326 (37.4)	
>3	391 (21.6)	292 (30.9)	100 (11.5)	
CD4 cell count category				<.001
<50 cells/mm ³	146 (8.1)	115 (12.3)	31 (3.6)	
50–199 cells/mm ³	256 (14.2)	151 (16.2)	105 (12.2)	
200–349 cell/mm ³	458 (25.3)	211 (22.5)	247 (28.4)	
350–499 cell/mm ³	514 (28.4)	257 (27.4)	257 (29.6)	
> = 500 cell/mm ³	434 (24.0)	205 (21.8)	229 (26.4)	
Missing	5 (0.2)	3 (0.3)	2 (0.2)	
Opportunistic infections	N = 1,049	N = 938	N = 111	
TB	117 (11.2)	88 (9.4)	29 (26.1)	<.001
PCP	15 (1.4)	15 (1.6)	0 (0)	.39
Candida	90 (8.5)	79 (8.4)	11 (9.9)	.59
Herpes Zoster	28 (2.4)	22 (2.4)	6 (5.4)	.04
Chronic Diarrhea	67 (6.4)	61 (6.5)	6 (5.5)	.84
Risk group, (n/N, %)	N = 1,727	N = 906	N = 821	
IDU	739 (42.8)	437 (48.2)	302 (36.8)	<.001
MSM	7 (0.4)	6 (0.7)	1 (0.1)	.13
FSW	9 (0.5)	5 (0.6)	4 (0.5)	1.00

IDU, injection drug users; FSW, female sex workers; MSM, men who have sex with men; PFR, proportion false recent.

Table 2

PFR for the LAg-Avidity EIA and BED-CEIA With and Without Low Viral Load Specimens and by Region, Vietnam 2009–2010

Method	Total false recents <i>n/N</i>	PFR% (95% CI)	Total false recents after excluding low viral load specimens <i>n/N</i>	PFR after exclusion% (95% CI)
BED-CEIA	41/1,813	2.3% (1.7–3.1) ^a	31/1,803	1.7% (1.2–2.4)
North	32/942	3.4% (2.4–4.8)	26/936	2.8% (1.9–4.0)
South	9/871	1.0% (0.5–2.0)	5/867	0.6% (0.2–1.3)
LAg-Avidity EIA	31/1,813	1.7% (1.2–2.4) ^a	22/1,791	1.2% (0.8–1.9)
North	25/942	2.7% (1.8–3.9)	19/936	2.0% (1.3–3.1)
South	6/871	0.7% (0.3–1.5)	3/868	0.3% (0.1–1.1)

^aMcNemar *p*-value .06.

95% confidence interval calculated using Wilson score interval procedure.

BED-CEIA, BED capture enzyme immunoassay.

Table 3

Proportion False Recent on the BED-CEIA Assay by Select Demographic and Clinical Characteristics and by Region, Vietnam 2009–2010

Characteristic	Total (n/N, %)	North (n/N, %)	South (n/N, %)	p-value
Proportion false recent	41/1,813 (2.3)	32/942 (3.4)	9/871 (1.0)	<.001
Age group (years)				.80
18–24	5/201 (2.5)	3/68 (4.4)	2/133 (1.5)	
25–34	33/1,182 (2.8)	26/585 (4.4)	7/597 (1.2)	
35–44	2/355 (0.6)	2/234 (0.9)	0/121 (0)	
45–54	1/67 (1.5)	1/50 (2.0)	0/17 (0)	
55+	0/8 (0)	0/5 (0.0)	0/3 (0)	
Sex				.02
Male	20/968 (2.1)	19/558 (3.4)	1/410 (0.2)	
Female	21/836 (2.5)	13/375 (3.5)	8/461 (1.7)	
Documented duration since last HIV positive test result (years)				.89
1–2	20/853 (2.3)	15/408 (3.7)	5/445 (1.1)	
2–3	9/569 (1.6)	7/243 (2.9)	2/326 (0.6)	
3+	12/391 (3.1)	10/291 (3.4)	2/100 (2.0)	
D4 cell count category				.05
<50 cells/mm ³	5/146 (3.4)	5/115 (4.3)	0/31 (0)	
50–199 cells/mm ³	3/256 (1.2)	1/151 (0.7)	2/105 (1.9)	
200–349 cell/mm ³	7/458 (1.5)	7/211 (3.3)	0/247 (0)	
350–499 cell/mm ³	14/514 (2.7)	12/257 (4.7)	2/257 (0.8)	
>= 500 cell/mm ³	12/434 (2.8)	7/205 (3.4)	5/229 (2.2)	
Opportunistic infections				NA
TB	2/117 (1.7)	2/88 (2.3)	0/29 (0)	
PCP	0/15 (0)	0/15 (0)	0/0 (0)	
Candida	4/90 (4.4)	4/79 (5.1)	0/11 (0)	
Herpes Zoster	0/28 (0)	0/22 (0)	0/6 (0)	
Chronic Diarrhea	1/67 (1.5)	1/61 (1.6)	0/6 (0)	
Risk group				.14
IDU	16/739 (2.2)	14/437 (3.2)	2/302 (0.7)	
MSM	0/7 (0)	0/6 (0)	0/1 (0)	
FSW	0/9 (0)	0/5 (0)	0/4 (0)	

NA, not applicable.

Table 4

Proportion False Recent on the LAg-Avidity EIA Assay by Select Demographic and Clinical Characteristics and by Region, Vietnam 2009–2010

Characteristic	Total (n/N, %)	North (n/N, %)	South (n/N, %)	p-value
Proportion false recent	31/1,813 (1.7)	25/942 (2.7)	6/871 (0.7)	.001
Age group (years)				1.00
18–24	4/201 (2.0)	3/68 (4.4)	1/133 (0.8)	
25–34	25/1,182 (2.1)	20/585 (3.4)	5/597 (0.8)	
35–44	1/355 (0.3)	1/234 (0.4)	0/121 (0)	
45–54	1/67 (1.5)	1/50 (2.0)	0/17 (0)	
55+	0/8 (0)	0/5 (0)	0/3 (0)	
Sex				.65
Male	15/968 (1.5)	13/558 (2.3)	2/410 (0.5)	
Female	16/836 (1.9)	12/375 (3.2)	4/461 (0.9)	
Documented duration since last HIV positive test result (years)				.69
1–2	17/853 (2.0)	13/408 (3.2)	4/445 (0.9)	
2–3	9/569 (1.6)	7/243 (2.9)	2/326 (0.6)	
3+	5/391 (1.3)	5/291 (1.7)	0/100 (0)	
CD4 cell count category				
<50 cells/mm ³	2/146 (1.4)	2/115 (1.7)	0/31 (0)	
50–199 cells/mm ³	2/256 (0.8)	2/151 (1.3)	0/105 (0)	.50
200–349 cell/mm ³	6/458 (1.3)	6/211 (2.8)	0/247 (0)	
350–499 cell/mm ³	10/514 (1.9)	8/257 (3.1)	2/257 (0.8)	
>= 500 cell/mm ³	11/434 (2.5)	7/205 (3.4)	4/229 (1.7)	
Opportunistic infections				NA
TB	1/117 (0.9)	1/88 (1.1)	0/29 (0)	
PCP	0/15 (0)	0/15 (0)	0/0 (0)	
Candida	1/90 (1.1)	1/79 (1.3)	0/11 (0)	
Herpes Zoster	0/28 (0)	0/22 (0)	0/6 (0)	
Chronic Diarrhea	0/67 (0)	0/61 (0)	0/6 (0)	
Risk group				
IDU	9/739 (1.2)	8/437 (1.8)	1/302 (0.3)	.63
MSM	0/7 (0)	0/6 (0)	0/1 (0)	
FSW	0/9 (0)	0/5 (0)	0/4 (0)	

NA, not applicable.